A CONJECTURE IN GAME THEORY

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We consider a team composed of n players, with each member playing the same r games, G_1 , G_2 , ..., G_r . We assume that each game G_j has two possible outcomes, success and failure, and that the probability of success in game G_j is equal to p_j for each player. We let X_{ij} be equal to one (1) if player i has a success in game j and let X_{ij} be equal to zero (0) if player i has a failure in game j. We assume throughout this paper that the random variables X_{ij} , $i = 1, 2, \ldots, n$, $j = 1, 2, \ldots, r$ are independent.

Let S_{jn} denote the total number of successes in the *j*th game. We define the point-value of a team to be

$$\Psi_n = \min_{1 \le j \le r} S_{jn} \, .$$

This means that the point-value of a team is equal to the minimum number of successes in any particular game. Clearly,

..., n,

$$P\{S_{jn} = m\} = {n \choose m} p_j^m (1 - p_j)^{n-m}, m = 0, 1, 2,$$

and

(1)

$$E[\Psi_n] = \sum_{k=0}^n k P\{\Psi_n = k\} = \sum_{k=0}^{n-1} P\{\Psi_n > k\}$$

= $\sum_{k=0}^{n-1} P\{S_{1n} > k, S_{2n} > k, \dots, S_{rn} > k\}$
= $\sum_{k=0}^{n-1} \prod_{j=1}^r P\{S_{jn} > k\}$
= $\sum_{k=0}^{n-1} \prod_{j=1}^r \sum_{m=k+1}^n \binom{n}{m} p_j^m (1 - p_j)^{n-m}.$

It follows from the definition of Ψ_n that the expected point-value for a team is an increasing function of n, i.e.,

 $E[\Psi_n] \leq E[\Psi_{n+1}], n = 1, 2, 3, \dots$

Since a team can add players in order to increase its expected point-value, it seems reasonable to define the score to be the expected point-value per player. Namely, we denote the score by

$$W_n = \frac{1}{n} E[\Psi_n].$$

Thus, from (1), we obtain

(2)
$$W_n = \frac{1}{n} \sum_{k=0}^{n-1} \prod_{j=1}^r \sum_{m=k+1}^n \binom{n}{m} p_j^m (1 - p_j)^{n-m}.$$

It is not obvious from (2) how the score varies as the number of players increases. We now prove that W_n is a strictly increasing function of n in the special case r = 2 and $p_1 = p_2$. We first prove three lemmas, which are also of independent interest.

Lemma 1: Let a team be composed of j players, with each member playing the same two games, G_1 and G_2 . Let the probability of success for each player in both games G_1 and G_2 be equal and be denoted by p. Let $u_j = P\{S_{1j} = S_{2j}\}$, for all positive integers j. Then

$$\frac{1}{2\Pi}\int_0^{2\Pi} |p + qe^{i\theta}|^{2j}d\theta = u_j,$$

where q = 1 - p.

Proof: Using the fact that

$$P\{S_{ij} = m\} = {\binom{j}{m}} p^m (1 - p)^{j - m}, m = 0, 1, 2, \dots, j, i = 1, 2,$$

and the independence of the random variables S_{1j} and S_{2j} , we obtain

(3)
$$u_{j} = \sum_{m=0}^{j} \left[\binom{j}{m} \right]^{2} p^{2m} (1-p)^{2(j-m)}, \ j = 1, 2, 3, \dots$$

We note that if f is the polynomial $f(z) = \sum_{m=0}^{J} \alpha_m z^m$, then

(4)
$$\frac{1}{2\Pi} \int_0^{2\Pi} |f(e^{i\theta})|^2 d\theta = \sum_{m=0}^J \alpha_m^2$$

We now apply the binomial expansion and (4) to the function $f(z) = (p + qz)^{j}$, where j is a positive integer. The binomial expansion yields

$$f(z) = (p + qz)^{j} = \sum_{m=0}^{j} \left[\binom{j}{m} p^{m} q^{j-m} \right] z^{j-m},$$

and using (3) and (4), we obtain

(5)
$$\frac{1}{2\Pi} \int_0^{2\Pi} |p + qe^{i\theta}|^{2j} d\theta = \sum_{m=0}^j \left[\binom{j}{m} \right]^2 p^{2m} q^{2(j-m)} = u_j.$$

Lemma 2: Let r = 2, $p_1 = p_2$, and $u_j = P\{S_{1j} = S_{2j}\}$, for all positive integers j. Then $u_j < u_{j-1}$.

and
$$\begin{aligned} |p + qe^{i\theta}|^2 \leq 1, \text{ for } 0 \leq \theta \leq 2\Pi\\ |p + qe^{i\theta}|^2 < 1, \text{ for } 0 < \theta < 2\Pi\end{aligned}$$

the desired result follows from (5).

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 $\underbrace{ \text{Lemma 3:}}_{\text{Let }d_j} = \mathbb{P}\{S_{1j} = S_{2j}\}, \text{ for all positive integers } j \text{ and let } u_0 = 1. \\ \underbrace{\text{Let }d_j}_{j} = \Psi_{j+1} - \Psi_j, \ j = 0, \ 1, \ 2, \ \dots, \text{ and let } \Psi_0 = 0.$ Then

(6)
$$E[d_j] = u_j p^2 + (1 - u_j)p$$
.

 $\underline{\textit{Proof}}$: Clearly, d_j can assume only the values 0 and 1 with the following probabilities:

$$P\{d_j = 0\} = 1 - [u_j p^2 + (1 - u_j)p],$$

$$P\{d_j = 1\} = u_j p^2 + (1 - u_j)p.$$

Since $E[d_j] = 0 \cdot P\{d_j = 0\} + 1 \cdot P\{d_j = 1\}$, we obtain the desired result.

<u>Theorem</u>: Let a team be composed of n players, with each member playing the same two games, G_1 and G_2 . Let the probability of success for each player in both games G_1 and G_2 be equal and be denoted by p. Then

$$W_n < W_{n+1}, n = 1, 2, 3, \ldots$$

Proof: Using the definition of W_n , we obtain

(7)
$$W_{n+1} - W_n = E\left[\frac{\Psi_{n+1}}{n+1} - \frac{\Psi_n}{n}\right] = \frac{1}{n(n+1)} E[n(\Psi_{n+1} - \Psi_n) - \Psi_n].$$

Using d_j , as defined in Lemma 3, and noting that $\Psi_n = \sum_{j=0}^{n-1} d_j$, (7) reduces to

$$W_{n+1} - W_n = \frac{1}{n(n+1)} E\left[nd_n - \sum_{j=0}^{n-1} d_j\right]$$

Using (6), we obtain

$$W_{n+1} - W_n = \frac{1}{n(n+1)} \left[n(u_n p^2 + (1 - u_n)p) - \sum_{j=0}^{n-1} (u_j p^2 + (1 - u_j)p) \right].$$

Thus, to prove that $W_n < W_{n+1}$, it suffices to show that

(8)
$$n(u_np^2 + (1 - u_n)p) - \sum_{j=0}^{n-1} (u_jp^2 + (1 - u_j)p) > 0.$$

Proving inequality (8) is equivalent to showing that

(9)
$$nu_n - \sum_{j=0}^{n-1} u_j = \sum_{j=1}^n j(u_j - u_{j-1}) < 0.$$

Since (9) follows from Lemma 2, we conclude that

 $W_n < W_{n+1}, n = 1, 2, 3, \ldots$

It is the author's conjecture that in the general case discussed in the beginning of this paper $(r > 2 \text{ and } p_1 \text{ not necessarily equal to } p_2)$ that W_n is a strictly increasing function of n, too. The above proven theorem and some elementary numerical computations suggest the truth of this statement, but the author has not been able to supply a complete proof.
